Role of Educational Exposure in the Association Between Myopia and Birth Order

Jeremy A. Guggenheim, PhD; Cathy Williams, MBBS, PhD, FRCOphth; for the UK Biobank Eye and Vision Consortium

IMPORTANCE Visual impairment due to myopia is an important public health issue. A prior analysis of population-based cohorts aged 15 to 22 years recruited from the United Kingdom and Israel suggested myopia and high myopia were approximately 10% more common in first-born compared with later-born children.

OBJECTIVE To examine whether myopia was associated with birth order in an earlier generation than studied previously and, if so, whether the association was attenuated after adjusting for education exposure, as predicted by the hypothesis that the education of children with later birth orders is less intense.

DESIGN, SETTING, AND PARTICIPANTS Cross-sectional study of UK Biobank participants recruited from 2006 to 2010. Analysis was restricted to participants aged 40 to 69 years who had a vision assessment, self-reported white ethnicity, and no history of eye disorders (N = 89,120). Myopia and high myopia were defined as autorefraction of −0.75 diopters (D) or less and −6.00 D or less, respectively.

EXPOSURES Birth order and information on potential confounders including highest educational qualification ascertained using a structured questionnaire.

MAIN OUTCOMES AND MEASURES Odds ratios (ORs) for myopia and high myopia by birth order, using logistic regression and adjusting for age and sex (model 1) or age, sex, and highest educational qualification (model 2).

RESULTS In model 1 (no adjustment for education), birth order was associated with both myopia and high myopia (eg, comparing first- vs second-born individuals; OR, 1.12; 95% CI, 1.08-1.16; P = 1.40E-11 and OR, 1.21; 95% CI, 1.11-1.30; P = 3.60E-06 for myopia and high myopia, respectively). The risk for myopia became progressively lower for later birth orders, suggesting a dose response. In model 2 (after adjusting for education), the effect sizes were attenuated by approximately 25% (OR, 1.09; 95% CI, 1.05-1.12; P = 1.30E-06 and OR, 1.15; 95% CI, 1.06-1.25; P = 4.60E-04 for myopia and high myopia, respectively) and the apparent dose response was abolished.

CONCLUSIONS AND RELEVANCE These data suggest that the association between birth order and myopia is not due to a new environmental pressure in the last 30 to 40 years. The attenuated effect size after adjusting for educational exposure supports a role for reduced parental investment in education of children with later birth orders in their relative protection from myopia.
Myopia is increasing in prevalence in younger generations in many parts of the world, and because the condition is a cause of visual impairment and blindness—either directly through myopic chorioteratal atrophy and choroidal neovascularization or indirectly through predisposition to cataract, glaucoma, and retinal detachment—it is becoming an increasingly important public health issue. Major known risk factors for myopia are genetic background, time spent outdoors, and time spent doing near work (including educational activities). However, refractive error is also associated with early life and life-course factors, such as maternal age, maternal smoking, gestational age, season of birth, and birth order.

In a study of 4 groups of participants (4401 children aged 15 years from a UK birth cohort; 888 277 Israeli Defense Force recruits aged 16-22 years; 1959 Singaporean children aged 13 years; and 1344 young adults aged 20 years from an Australian birth cohort), there was strong statistical support for an association between birth order and myopia in the 2 larger samples; however, there was weak/little support for the 2 smaller samples. In the largest cohort, there was also evidence for a dose-response relationship, with myopia being increasingly less common in individuals the more older siblings they had. One potential cause of the association between birth order and myopia is parental investment in education. On average, parents have been reported to direct more of their available resources to earlier-born children, resulting in better educational attainment in earlier-born than later-born individuals. Thus, parents may expose their earlier-born children to a more myopia-predisposing environment. Here, we sought to replicate the previously reported association between birth order and myopia in an older sample of UK adults and to examine whether adjusting for educational exposure attenuates any association observed in this sample.

Methods

UK Biobank Assessments

The UK Biobank recruited 502 649 participants aged 37 to 73 years from 2006 to 2010. Participants attended 1 of 22 assessment centers, at which they completed a touch-key questionnaire, had a face-to-face interview with a trained nurse, and underwent physical assessments. During later stages of recruitment, the assessments included an ophthalmic component. All assessments adhered to standardized protocols. Ethical approval was obtained from the National Health Service National Research Ethics Service (Ref 11/NW/0382) and all participants provided written informed consent.

The touch-key questionnaire included information about demographics and potential confounders, namely: race/ethnicity, ophthalmic history, number of total/older siblings, birth weight, maternal age, current time spent outdoors during summer (“In a typical day in summer, how many hours do you spend outdoors?”), and educational or professional qualifications (“Which of the following qualifications do you have [you can select more than one]?”; with the options, “College or university degree, A-levels/AS-levels, O-levels, [Certificate of Secondary Education] or equivalent”; “[National Vocational Qualification] or [Higher National Diploma] or [Higher National Certificate] or equivalent”; or “other professional qualifications, eg: nursing, teaching, none of the above”). Age at which continuous full-time education was completed was asked of individuals not reporting a college or university degree. Refractive error was measured by noncycloplegic autorefraction (Tomey RC5000 autorefractor) after removing habitual spectacles or contact lenses, as part of the ophthalmic assessment.

Classification of Participant Demographics and Ocular Phenotype

Birth orders of 4 and greater were combined owing to small numbers. Nonsingletons (eg, twins) were excluded. Race/ethnicity was classified as either white (self-report of British, Irish, or any other white background) or other (self-report of Indian, Pakistani, African, Chinese, mixed race, or “prefer not to answer”). Individuals who reported nonwhite race/ethnicity and those aged younger than 40 years or older than 69 years were excluded owing to their low numbers, especially for higher birth orders. Because the relationship between age and the prevalence of myopia was nonlinear (Figure), age was modeled as a categorical variable in 3-year intervals (40-42, 43-45, 46-48, 49-51, 52-54, 55-57, 58-60, 61-63, 64-66, and 67-69 years). Maternal age was categorized into 5 groups (<20, 20-24, 25-29, 30-34, and >34 years). Self-reported birth weight was filtered to exclude participants with a z score greater than 4. Each participant was assigned a Townsend Deprivation Index score corresponding to their postcode area, based on the preceding national census output areas. Highest educational qualification was categorized into 4 levels: none; O-levels or Certificate of Secondary Education or equivalent; A-levels/AS-levels, National Vocational Qualification, Higher National Diploma, or Higher National Certificate or equivalent or other professional qualification; and degree. For the Biobank participants’ generation, the UK school system provided free universal compulsory education between the ages of 5 and 15 to 16 years. Standard examinations were taken at the ages of 16 (O-levels and Certificate of Secondary Education).
Education), 17 (A/S-levels), and 18 (A-levels) years. National Vocational Qualification, Higher National Diploma, and Higher National Certificate refer to vocational qualifications that required approximately 2 or more years of part-time or full-time study after the age of 15 years.

Participants were excluded if they reported a history of cataract, cataract surgery, corneal graft surgery, laser eye surgery, or serious eye trauma, as were autorefraction readings if accompanied by a low reliability or lower reliability error message. Spherical equivalent was calculated as the spherical power plus half the cylinder power and averaged between fellow eyes. Individuals with a refractive error of \(-0.75\) diopters (D) or less and \(-6.00\) D or less were classified as having myopia and high myopia, respectively.

**Statistical Analyses**

The odds ratio (OR) for myopia in participants of birth orders 1 through 4 or more was estimated using logistic regression. An initial analysis was conducted that adjusted for age and sex only (model 1) followed by analyses that also included highest educational qualification (model 2) and highest educational qualification, maternal age, birth weight, Townsend Deprivation Index score (natural-log-transformed to remove skew), and time currently spent outdoors in summer (model 3). Analogous models were used to calculate the OR for high myopia vs nonmyopia (this resulted in a reduction in participant numbers owing to the exclusion of mild/moderate myopes). A final model was used to gauge an alternative measure of educational exposure; this model adjusted for age, sex, and age at completion of full-time education (model 4). Details of participants included and excluded from each analysis model are presented in eTable 1, eTable 2, and eTable 3 in the Supplement. An analysis of the relationship between birth order and the level of refractive error was also carried out (eAppendix in the Supplement). Individuals with no siblings (only children) were included in all analyses, except where indicated.

**Results**

Approximately 23% of UK Biobank participants underwent autorefraction and 89,120 were included in the analysis (white; aged 40-69 years; Table 1). The prevalence of myopia varied nonlinearly with age (Figure, A), while the number of siblings was relatively stable at an average of 2 (Figure, B).

**Birth Order and Myopia**

In analyses adjusted for age and sex, the OR for myopia was 1.12 (95% CI, 1.08-1.16) for first-born vs second-born individuals, and this increased to 1.38 (95% CI, 1.31-1.46) for first-born vs fourth- or higher-born individuals. The corresponding ORs obtained after including highest educational qualification were reduced, especially for higher birth orders (Table 2; model 2). Thus, for first-born vs second-born individuals, the OR reduced by 25% from 1.12 to 1.09, while for first-born vs fourth- or higher-born individuals, the OR reduced by almost 50% from 1.38 to 1.17.

Further analyses were carried out in a subset of participants (n = 25,278) with data available for a range of potential confounders. In this subset, there was a smaller association between birth order and myopia than was observed in the full sample (model 1 results in Table 2 vs Table 3). Adjusting for highest educational qualification attenuated these associations further (Table 3; model 2), while adjusting for maternal age, social deprivation, birth weight, time spent outdoors currently in summer, and highest educational qualification yielded results comparable with the unadjusted analyses (Table 3; model 3).

The suggestion of a dose-response relationship between birth order and myopia, whereby myopia risk decreased with the number of older siblings, was substantially weakened or lost completely after adjusting for highest educational attainment (model 1 vs models 2 or 3) both for the full set of participants (Table 2) and the subset with full information (Table 3).
Birth Order and High Myopia

The relationship between birth order and high myopia (Table 2 for full sample and Table 3 for subset) shared several of the features of its relationship with myopia. In the full sample, after adjusting for age and sex (model 1), there was evidence for a relationship between high myopia and birth order, with an OR of 1.21 (95% CI, 1.11-1.30; n = 65 500) for first-vs-second-born individuals. However, there was no suggestion of a dose-response relationship between later birth order and

<table>
<thead>
<tr>
<th>Variable</th>
<th>Sample With Incomplete Information (n = 63 842)</th>
<th>Sample With Full Information (n = 25 278)</th>
<th>P Value</th>
</tr>
</thead>
<tbody>
<tr>
<td>Age, mean (SD), y</td>
<td>59.0 (7.2)</td>
<td>51.7 (7.2)</td>
<td>&lt;1.0E-99</td>
</tr>
<tr>
<td>Birth weight, mean (SD), kg</td>
<td>3.3 (0.7)</td>
<td>3.4 (0.6)</td>
<td>3.70E-07</td>
</tr>
<tr>
<td>Townsend Deprivation Index score, natural log, mean (SD)</td>
<td>2.0 (0.4)</td>
<td>2.0 (0.4)</td>
<td>7.60E-03</td>
</tr>
<tr>
<td>Time spent outdoors currently, mean (SD), h/d</td>
<td>3.9 (2.3)</td>
<td>3.4 (2.2)</td>
<td>&lt;1.0E-99</td>
</tr>
<tr>
<td>Age at completion of full-time education, mean (SD), ya</td>
<td>16.6 (2.1)</td>
<td>17.1 (2.0)</td>
<td>&lt;1.0E-99</td>
</tr>
<tr>
<td>Female</td>
<td>33 075 (51.8)</td>
<td>15 104 (59.8)</td>
<td>&lt;1.0E-99</td>
</tr>
<tr>
<td>Maternal age category, yb</td>
<td>&lt;20</td>
<td>383 (3.7)</td>
<td>888 (3.5)</td>
</tr>
<tr>
<td>20-24</td>
<td>2842 (27.8)</td>
<td>7842 (31.0)</td>
<td></td>
</tr>
<tr>
<td>25-29</td>
<td>3987 (38.9)</td>
<td>9894 (39.1)</td>
<td>4.00E-14</td>
</tr>
<tr>
<td>30-34</td>
<td>2149 (21.0)</td>
<td>4955 (19.6)</td>
<td></td>
</tr>
<tr>
<td>≥35</td>
<td>879 (8.6)</td>
<td>1699 (6.7)</td>
<td></td>
</tr>
<tr>
<td>Birth orderb</td>
<td>First born</td>
<td>28 598 (44.8)</td>
<td>13 621 (53.9)</td>
</tr>
<tr>
<td></td>
<td>Second born</td>
<td>19 457 (30.5)</td>
<td>7583 (30.0)</td>
</tr>
<tr>
<td></td>
<td>Third born</td>
<td>8432 (13.2)</td>
<td>2738 (10.8)</td>
</tr>
<tr>
<td></td>
<td>Fourth or higher born</td>
<td>7355 (11.5)</td>
<td>1336 (5.3)</td>
</tr>
<tr>
<td>Highest educational qualificationb</td>
<td>None</td>
<td>11 375 (17.8)</td>
<td>1693 (6.7)</td>
</tr>
<tr>
<td></td>
<td>O-levels, CSEs, or equivalent</td>
<td>16 575 (26.0)</td>
<td>7496 (29.7)</td>
</tr>
<tr>
<td></td>
<td>A-levels, professional, or equivalent</td>
<td>14 891 (23.3)</td>
<td>5708 (22.6)</td>
</tr>
<tr>
<td></td>
<td>Degree</td>
<td>21 001 (32.9)</td>
<td>10 381 (41.1)</td>
</tr>
<tr>
<td>Refractive error categoryb</td>
<td>Nonmyopic, &gt;−0.75 D</td>
<td>45 059 (70.6)</td>
<td>16 898 (66.8)</td>
</tr>
<tr>
<td></td>
<td>Low/moderate myopia, ≤−0.75 D and &gt;−6.00 D</td>
<td>16 389 (25.7)</td>
<td>7231 (28.6)</td>
</tr>
<tr>
<td></td>
<td>High myopia, ≤−6.00 D</td>
<td>2394 (3.7)</td>
<td>1149 (4.5)</td>
</tr>
</tbody>
</table>

Birth Order and Refractive Error

After adjusting for age and sex, there was evidence suggesting a dose-response relationship between later birth order and
namely, including only families containing 2 children. There
ses using a statistical approach to control for family size,13,14
myopiacoobeafeatureofamoregeneralunderlyingasso-
to investigate whether the association between birth order and
and 46% second born compared with the 50%-50% propor-
ten years old, we also carried out analyses adjusted for age at
on of near work undertaken. Therefore, we also carried out
cluded.13,14 For high myopia (prevalence of approximately 30%
ies were 30 727 participants who had only 1 sibling (54% first born
and myopia could be a feature of a more general underlying asso-
cannot capture all aspects of the myopia-predisposing influ-
et sequence; evidence from the UK Biobank suggests that high
were again similar to those observed when adjusting for highest
for age and sex only.
† Sample size reduced because participants with mild/moderate
the 2015 American Medical Association. All rights reserved.
lecular age. Adjusting for age and sex yielded an OR for myopia in
of 1.08 (95% CI, 1.03-1.13; P = 1.5E-03). These estimates were
with our previous analyses, using this measure of educational
mended expected under random ascertainment). Adjusting for age
y had very little effect, thus confirming that only children were
the Supplement). Adjusting for Age at Completion of Full-Time
and highest educational qualification.‡ Sample size reduced be-
Table 4. ORs for Myopia and High Myopia by Birth Order After Adjusting for the Full Set of Potential Confounders

<table>
<thead>
<tr>
<th>Variable</th>
<th>Model 1a OR (95% CI)</th>
<th>P Value</th>
<th>Model 2b OR (95% CI)</th>
<th>P Value</th>
<th>Model 3c OR (95% CI)</th>
<th>P Value</th>
</tr>
</thead>
<tbody>
<tr>
<td>Myopia (n = 25 278)</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>First vs second born</td>
<td>1.10 (1.04-1.17)</td>
<td>1.70E-03</td>
<td>1.08 (1.01-1.15)</td>
<td>1.60E-02</td>
<td>1.14 (1.07-1.21)</td>
<td>6.10E-05</td>
</tr>
<tr>
<td>First vs third born</td>
<td>1.09 (1.00-1.19)</td>
<td>6.00E-02</td>
<td>1.04 (0.95-1.14)</td>
<td>3.60E-01</td>
<td>1.15 (1.05-1.27)</td>
<td>3.20E-03</td>
</tr>
<tr>
<td>First vs fourth born or higher</td>
<td>1.16 (1.02-1.30)</td>
<td>1.90E-02</td>
<td>1.02 (0.90-1.15)</td>
<td>7.50E-01</td>
<td>1.17 (1.03-1.34)</td>
<td>1.80E-02</td>
</tr>
<tr>
<td>High myopia (n = 18 047d)</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>First vs second born</td>
<td>1.18 (1.03-1.36)</td>
<td>2.10E-02</td>
<td>1.14 (0.99-1.31)</td>
<td>7.30E-02</td>
<td>1.23 (1.06-1.43)</td>
<td>5.10E-03</td>
</tr>
<tr>
<td>First vs third born</td>
<td>1.18 (0.96-1.44)</td>
<td>1.20E-01</td>
<td>1.11 (0.90-1.36)</td>
<td>3.30E-01</td>
<td>1.29 (1.04-1.60)</td>
<td>2.20E-02</td>
</tr>
<tr>
<td>First vs fourth born or higher</td>
<td>0.93 (0.72-1.19)</td>
<td>5.50E-01</td>
<td>0.77 (0.60-1.00)</td>
<td>4.70E-02</td>
<td>0.95 (0.73-1.25)</td>
<td>7.40E-01</td>
</tr>
</tbody>
</table>

Abbreviation: OR, odds ratio.
* Model 1 adjusted for age and sex only.
† Model 2 adjusted for age, sex, and maternal age.
‡ Model 3 adjusted for age, sex, highest educational qualification, Townsend Deprivation Index score, current time spent outdoors, birth weight, and maternal age.
§ Sample size reduced because participants with mild/moderate myopia were excluded.

Table 4. ORs for Myopia and High Myopia by Birth Order After Adjusting for Age at Completion of Full-Time Education

<table>
<thead>
<tr>
<th>Variable</th>
<th>Model 1a OR (95% CI)</th>
<th>P Value</th>
<th>Model 4b OR (95% CI)</th>
<th>P Value</th>
</tr>
</thead>
<tbody>
<tr>
<td>Myopia (n = 57 447)</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>First vs second born</td>
<td>1.13 (1.08-1.18)</td>
<td>1.00E-07</td>
<td>1.09 (1.04-1.14)</td>
<td>1.50E-04</td>
</tr>
<tr>
<td>First vs third born</td>
<td>1.19 (1.12-1.27)</td>
<td>8.60E-09</td>
<td>1.11 (1.04-1.18)</td>
<td>8.40E-04</td>
</tr>
<tr>
<td>First vs fourth born or higher</td>
<td>1.31 (1.22-1.39)</td>
<td>3.50E-16</td>
<td>1.14 (1.07-1.22)</td>
<td>6.90E-05</td>
</tr>
<tr>
<td>High myopia (n = 44 473c)</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>First vs second born</td>
<td>1.21 (1.08-1.36)</td>
<td>1.60E-03</td>
<td>1.15 (1.03-1.30)</td>
<td>1.80E-02</td>
</tr>
<tr>
<td>First vs third born</td>
<td>1.22 (1.04-1.43)</td>
<td>1.40E-02</td>
<td>1.10 (0.93-1.29)</td>
<td>2.60E-01</td>
</tr>
<tr>
<td>First vs fourth born or higher</td>
<td>1.24 (1.05-1.47)</td>
<td>1.00E-02</td>
<td>1.02 (0.86-1.20)</td>
<td>8.40E-01</td>
</tr>
</tbody>
</table>

Abbreviation: OR, odds ratio.
* Model 1 adjusted for age and sex only.
† Model 2 adjusted for age, sex, and maternal age.
‡ Model 3 adjusted for age, sex, highest educational qualification, Townsend Deprivation Index score, current time spent outdoors, birth weight, and maternal age.
§ Sample size reduced because participants with mild/moderate myopia were excluded.

Educational Exposure, Myopia, and Birth Order

a more positive (residual) refractive error, which was again much reduced after adjusting for highest educational attainment (eFigure in the Supplement).

Adjusting for Age at Completion of Full-Time Education

Highest educational qualification may not capture all aspects of the myopia-predisposing influence of education, such as the amount of near work undertaken. Therefore, we also carried out analyses adjusted for age at completion of full-time education (n = 57 447; Table 4). In keeping with our previous analyses, using this measure of educational exposure led to a 31% attenuation of the association between birth order and myopia (prior to adjustment: OR, 1.13; after adjustment: OR, 1.09) and eliminated all evidence of a dose-response relationship (Table 4). In the case of birth order vs high myopia, adjusting for age at completion of full-time education also yielded results that were very similar to those observed when adjusting for highest educational qualification.

Family Size

To investigate whether the association between birth order and myopia could be a feature of a more general underlying association between family size and myopia, we repeated our analyses using a statistical approach to control for family size,13,14 namely, including only families containing 2 children. There were 30 727 participants who had only 1 sibling (54% first born and 46% second born compared with the 50%-50% proportions expected under random ascertainment). Adjusting for age and sex yielded an OR for myopia in first-born vs second-born individuals of 1.12 (95% CI, 1.06-1.17; P = 1.0E-05), while adjusting for age, sex, and highest educational qualification yielded an OR for myopia of 1.08 (95% CI, 1.03-1.13; P = 2.1E-03). These estimates were close to those for the full sample, suggesting that they were not driven by family size per se (Table 2).

Exclusion of Only Children

A proportion of first-born children will be only children, ie, individuals with no brothers or sisters. As shown in eTable 4 in the Supplement, repeating our analyses after excluding only children had very little effect, thus confirming that only children were not driving the associations.

Risk Ratios vs Odds Ratios

The prevalence of myopia in the UK Biobank sample was approximately 30% (Table 1). For such a highly prevalent condition, an OR will accentuate the true relative risk (RR); for example, ORs of 1.10 and 1.20 would correspond to RRs of 1.07 and 1.13, respectively.19 For high myopia (prevalence of approximately 4% among UK Biobank participants), the corresponding RRs would be 1.10 and 1.19, respectively. The (un-
Discussion

We observed strong evidence that first-born individuals were more often myopic than nonfirst-born UK Biobank participants, confirming previous findings. The magnitude of this association was small: first-born participants were approximately 10% more likely to be myopic than nonfirst-born participants, which equated to first-born individuals having a refractive error that was less than −0.25 D more negative, on average, than nonfirst-born participants. Much larger shifts toward a more negative refractive error have been observed in East and Southeast Asia over the past few decades, implicating additional environmental influences over those assessed in these analyses.

The results did not support the idea that the association between birth order and myopia arose through confounding via the participant demographic-related effects, age, sex, and socioeconomic status, nor the maternal/birth-related effects, maternal age, and birth weight. In contrast, there was evidence of confounding due to educational exposure. After adjusting for either of 2 measures of educational exposure—highest educational qualification or age at completion of full-time education—the association between birth order and myopia was attenuated and no dose-response relationship was evident. Morgan and Cotch suggested that such confounding was a plausible cause of the association between birth order and myopia based on reports that after controlling for family size, children with an earlier birth order do relatively better at school owing to parents investing more time, effort, and/or resources in educating children with an earlier birth order.

Greater educational exposure in earlier-born children may expose them to a more myopigenic environment; for example, more time doing near work and less time spent outdoors. Our findings that statistical adjustment for indices of educational exposure partially attenuated the magnitude of the association between birth order and myopia, and completely removed the evidence for a dose-response relationship, therefore support the idea that reduced parental investment in children’s education for offspring of later birth order contributed to the observed birth order vs myopia association and produced the observed dose-response relationship. However, because the increased risk for myopia in first-born vs non-first-born individuals was reduced but not abolished, either the statistical adjustment failed to adequately capture the true relationship fully or other unmeasured factor(s) contributed to the higher prevalence of myopia in first-born vs nonfirst-born individuals. These results add to the extensive literature implicating a role for education in the etiology of myopia, although a causal relationship cannot be confirmed using observational data.

The association between birth order and high myopia was similar in magnitude to that between birth order and any myopia and also was reduced by adjusting for educational exposure. This implies a role for environment (ie, education) in the etiology of high myopia as well as myopia.

The magnitude of the association between birth order and myopia appeared weaker in the subset with full data than the full sample (compare model 1 in Table 2 vs Table 3), especially as regards the dose-response relationship. A comparison of demographic characteristics between those with complete or incomplete information for potential confounders highlighted many differences, including age, sex, and highest educational attainment (Table 1). The lack of uniformity of the association in 2 subgroups of the sample argues against a biological factor, such as a parity-related maternal effect during pregnancy, being fully responsible for causing the association between birth order and myopia.

The strengths of this study were highly standardized methods of data collection, a large sample size, use of an objective and reliable method of quantifying refractive error in this age group, participant selection not being directly aimed at ocular health (thus reducing selection bias), and availability of information on a range of potential confounders. The weaknesses were using self-report to exclude participants with cataracts and the wide age range of the sample, which increased the risk for bias due to confounding between myopia and changing demographic variables. The 2 measures of education that were available may not have captured all relevant aspects of the educational process. Additionally, the participants were not selected at random from the population and had nonrandom variations in levels of missing information for covariates (Table 1); therefore, the results may not be fully representative of the general population. Finally, information on the time UK Biobank participants spent outdoors during childhood was not collected; therefore, any potential role of this important exposure in mediating the association of birth order and myopia could not be investigated.

Conclusions

First-born individuals in a sample of UK adults were approximately 10% more likely to be myopic or highly myopic than later-born individuals. The results replicate earlier findings from 2 contemporary international cohorts of adolescents/young adults, implying that the cause of the birth order–myopia association is widespread and has been in existence for several decades. The association was larger before adjusting for educational exposure, suggesting that reduced parental investment in the education of children of later birth order may be partly responsible.
Acquisition, analysis, or interpretation of data: Both authors.

Drafting of the manuscript: Both authors.

Critical revision of the manuscript for important intellectual content: Williams.

Statistical analysis: Both authors.

Obtained funding: Both authors.

Administrative, technical, or material support: Williams.

Conflict of Interest Disclosures: All authors have completed and submitted the ICMJE Form for Disclosure of Potential Conflicts of Interest and none were reported.

Funding/Support: The work was funded by a National Institute for Health Research Career Development Fellowship (CDF-2009-02-35; Dr Williams) and an internal award (ZOGM) from the Hong Kong Polytechnic University (Dr Guggenheim).

Role of the Funder/Sponsor: The funders had no role in the design and conduct of the study; collection, management, analysis, and interpretation of the data; preparation, review, or approval of the manuscript; and decision to submit the manuscript for publication.

Group Information: The UK Biobank Eye and Vision Consortium includes the following investigators: Cardiff University, John Gallacher, Jeremy Guggenheim, and James Morgan (steering committee member); City University (London), David Crabb and Haogang Zhu; Edinburgh University, Bal Dhillon (steering committee member); Gloucestershire Hospitals NHS Foundation Trust, Irene Stratton; King’s College London, Chris Hammond, Eoin O’Sullivan, and Katie Williams; Kingston University, Sarah Barman; Leeds Teaching Hospitals NHS Trust, Martin McIvor; Manchester University, Tariq Aslam and Paul Bishop (steering committee member); Moorfields Eye Hospital (London), Peter Blosy, Catey Bunce, Michelle Chan, Alexander Day, Parul Desai, Anthony Khawaja, Gerasimos Lascaratos, Praveen Patel, Tunde Peto, Nicholas Stoutthidis, Dhanes Thomas, Adnan Tufail, and Ananth Viswanathan; Newcastle University, David Steel; Queen’s University Belfast, Michelle McGaughie, Bernadette McGuinness, Gareth McKay, Usha Chakravarty, Ruth Hogg, and Anne Hughes; St George’s University of London, Chris Owen and Alicja Rudnicka; St James’s University Hospital (Leeds), Sarah Mackie; UCL Institute of Child Health, Philippa Cumberland and Jugnoo Rahi; UCL Institute of Ophthalmology, Antonietta Chianca, Valentina Cipriani, Paul Foster, David (Ted) Garway-Heath, Priyal Gupta, Pearse Keane, Peng Tey Khaw (steering committee member), Phil Luthert, Tony Moore, Zaynah Muthy, and Caroline Thaung; University Hospital, Nottingham, Stephen Vernon; University of Bristol, Andrew Dick (steering committee cochair), and Cathy Williams; University of Cambridge, John Yates, Jennifer Yip, and Keith Martin; University of Dundee, Emanuele Trucco; University of East Anglia, Carlota Grossi Sampedro and Max Yates; University of Edinburgh, Tom MacCulllary, Danny Mitry, and Cathie Sudlow; University of Essex, Yanchun Bao; University of Liverpool, Simon Harding (steering committee member), and University of Southampton, Srini Govender and Andrew Lotery (steering committee chair).

Additional Information: This research was conducted using the UK Biobank Resource (https://www.ukbiobank.ac.uk/).

REFERENCES


16. Grant RL. Converting an odds ratio to a range of plausible relative risks for better communication of research findings. BMJ. 2014;348:g4750.


