In Reply: We agree with Drs Rao and Dlouhy that fireworks do cause many serious eye injuries, especially among youth. Based on their comments, there appears to be a misunderstanding of what was done in our study. Previous research has shown that projectiles can cause serious eye injuries due to considerable deformation, penetration, or perforation of the eye. Our experiments were designed to evaluate overpressure (ie, pressure wave) from a fireworks explosion as a potential eye injury mechanism.

Following exposure to an overpressure event, eyes were visually examined for corneal abrasions, scleral damage, and globe rupture. The main observation in this study was that the overpressure caused by fireworks was not large enough to cause severe eye injuries. Specifically, no scleral damage or globe ruptures were observed on direct inspection, no considerable deformation was seen in the high-speed video, and the calculated injury risk for physiological eye injuries was less than or equal to 0.01%. Therefore, we suggested that clinically observed eye injuries from fireworks are caused by projectile impacts to the eye. This would include debris and fragmented material as well as aerial devices such as bottle rockets.

As noted in our article, physiological injuries such as scleral bleeding cannot be assessed using cadaver eyes. For ethical reasons, researchers cannot expose humans to the injurious conditions needed to obtain the data required to develop risk curves for physiological injuries. Therefore, intraocular pressure was used to calculate the risk of physiological injuries (hyphema, lens damage, and retinal damage) using published injury risk curves.

The referenced injury risk curves are based on in vivo animal testing. Specifically, 165 individual in vivo impact tests on monkey, pig, and cat eyes were used to develop the risk functions for hyphema, lens damage, and retinal damage. Additionally, 81 human cadaver eyes were tested to develop an injury risk function for globe rupture. These methods are commonly used in the field of injury biomechanics to develop injury risk curves and injury thresholds that can be used to predict human injury under various loading conditions. Although not developed for blast overpressure, these risk curves provide a reasonable assessment of potential injuries that cannot otherwise be directly assessed in a cadaver model.

Overall, our study serves as a first step in assessing blast overpressure as a potential eye injury mechanism. Future studies could evaluate other potential injuries such as lens damage and retinal tears in a cadaver model. We agree that the effect of higher overpressure on the eye should be investigated to more thoroughly evaluate overpressure as a potential eye injury mechanism.

However, based on the results of this and previous studies, the overpressure required to cause severe ocular injuries would likely result in life-threatening injuries that would take precedence over potential eye injuries. Additionally, overpressures of this magnitude would be more representative of blasts observed during military combat than from commercial fireworks.

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Conflict of Interest Disclosures: The authors have completed and submitted the ICMJE Form for Disclosure of Potential Conflicts of Interest and none were reported.


RESEARCH LETTER

Pseudoephedrine Sales and Seizures of Clandestine Methamphetamine Laboratories in Kentucky

To the Editor: The illicit production of methamphetamine from the precursor pseudoephedrine in clandestine laboratories fuels up to 35% of the domestic supply. Public health, law enforcement, and medical associations support restricted access to methamphetamine precursors; manufacturers oppose restrictions.

Kentucky law limits pseudoephedrine sales in all counties to 7.2 g/person/month (as of July 2012, formerly 9 g/person/month), sufficient to allow a patient to take the maximum daily dose (240 mg/d) each day. Electronic tracking of sales is also required. Despite these restrictions, increases in the number of reported methamphetamine labo-
atory seizures (called laboratories) continue.2–4 We analyzed the relationship between pseudoephedrine sales and the number of laboratories reported in Kentucky.

Methods. Using county level data from 2010, we examined the association between the natural log of the number of laboratories and county pseudoephedrine sales (grams/100 residents) adjusting for the number of law enforcement officers and the percentage of high school graduates. Law enforcement regulations define laboratories as “requiring two or more chemicals or two or more pieces of equipment used in manufacturing methamphetamine.”5

Pseudoephedrine sales data were obtained from the National Precursor Log Exchange, the real-time electronic system mandated for use by pharmacies to track all sales of nonprescription pseudoephedrine medications in Kentucky. The number of full-time law enforcement officers in a county was used as a proxy for the influence of law enforcement on laboratory seizures. Data on laboratories and law enforcement officers were obtained from the Kentucky State Police. The percentage of high school graduates was used as a socioeconomic control and was obtained from the US Census.

We first visualized the relationship using a scatterplot, and then calculated a log-linear regression model using Stata version 11 (StataCorp). All reported $P$ values are 2-sided, with $P < 0.05$ considered statistically significant. The University of Kentucky institutional review board approved the study.

Results. Four Kentucky counties reported no pseudoephedrine sales, leaving 116 counties for analysis. In 2010, Kentuckians purchased a mean 24,664 g (SD, 60,313 g) of pseudoephedrine per county (per county mean, 49 [SD, 39] g/100 residents) and there were 1072 laboratories reported (per county mean, 9.28 [SD, 20.91] laboratories). There was considerable variability in pseudoephedrine sales per county from 0.26 g/100 residents (population: 13,752) to 147 g/100 residents (population: 25,581).

Variability in the number of laboratories was also evident from 0 in many counties to 134 laboratories. The natural log of the number of laboratories reported in each county and the grams of pseudoephedrine sold per 100 residents (FIGURE) indicates a relationship between sales and laboratories. Regression results in the TABLE were consistent with the visual representation in the Figure.

Counties with greater pseudoephedrine sales were significantly associated with greater numbers of laboratories. In this model, a 1-g increase in pseudoephedrine sales per 100 people was associated with a 1.7% increase in laboratories. For a typical county, a 13-g per 100 resident increase in pseudoephedrine sales was associated with approximately 1 additional laboratory.

Comment. The strength of this study is that it is the first, to our knowledge, to provide empirical evidence that pseudoephedrine sales are correlated with the clandestine manufacture of methamphetamine. While the incidence of conditions for which pseudoephedrine is indicated is not known, and may vary by county, our results indicated a 365-fold variation in pseudoephedrine sales between Kentucky counties.

Our study is limited by the ecological design, possible underdetection and underreporting of laboratories, purchase of pseudoephedrine across county lines, and importation of pseudoephedrine into Kentucky. In addition, law enforcement’s use of pseudoephedrine sales data to identify questionable pseudoephedrine purchases could have affected the association between pseudoephedrine sales and laboratories. Nevertheless, this study highlights the need for research on various approaches to containing clandestine methamphetamine production, including restriction of pseudo-

Table. Statistical Modeling of Methamphetamine Laboratory Seizures in 116 Kentucky Counties in 2010

<table>
<thead>
<tr>
<th></th>
<th>Bivariate Model</th>
<th>Multivariate Model</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>Mean (SD)</td>
<td>Coefficient (SE)</td>
</tr>
<tr>
<td>Pseudoephedrine sales, g/100 county residents</td>
<td>49 (38)</td>
<td>0.015 (0.003)</td>
</tr>
<tr>
<td>No. of full-time law enforcement officers/100 county residents</td>
<td>0.12 (0.11)</td>
<td>NA</td>
</tr>
<tr>
<td>Adult population with a high school diploma, %</td>
<td>0.76 (0.08)</td>
<td>NA</td>
</tr>
</tbody>
</table>

Abbreviation: NA, not applicable.

*Although other models were used to fit the data, including Poisson, zero-inflated Poisson, and negative binomial, the residual distribution indicated an ordinary least-squares linear regression model on the natural log of the number of laboratories in a county as the most appropriate. No pseudoephedrine sales were reported for 4 counties (Bracken, Hancock, Owsley, and Robertson) and they were excluded from the analysis.
doephedrine sales to only those patients who have a true medical need for its decongestant properties.

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Author Contributions: Dr Talbert had full access to all of the data in the study and takes responsibility for the integrity of the data and the accuracy of the data analysis.

Study concept and design: Talbert, Blumenschein, Stromberg, Freeman.

Acquisition of data: Talbert, Freeman.

Analysis and interpretation of data: Talbert, Blumenschein, Burke, Stromberg, Freeman.

Drafting of the manuscript: Talbert, Blumenschein, Stromberg, Freeman.

Critical revision of the manuscript for important intellectual content: Talbert, Blumenschein, Burke, Stromberg, Freeman.

Statistical analysis: Talbert, Burke, Stromberg.

Administrative, technical, or material support: Talbert.

Study supervision: Talbert, Stromberg.

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Disclaimer: The content is solely the responsibility of the authors and does not necessarily represent the official views of the National Institutes of Health.

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CORRECTION

Incorrect Author’s Last Name and Errors in the Text: In the Original Contribution titled “Effect of Biolimus-Eluting Stents With Biodegradable Polymer vs Bare-Metal Stents on Cardiovascular Events Among Patients With Acute Myocardial Infarction: The COMFORTABLE AMI Randomized Trial” by Raber et al, published in the August 22/29, 2012, issue of JAMA (2012;308[8]:777-787), one of the names in the byline had an incorrect last name. The name should have appeared as Miodrag Ostojic, MD. This author’s last name was also misspelled in the Author Affiliations and Author Contributions sections. In the second paragraph of the Methods, Procedures section, 14 mm was omitted from the lengths of stent types. In Table 1, “No./total No. (%)” should have been deleted from the Killip class II, III, or IV row of the table. In the second paragraph of the Comment section, “4.9% absolute reduction” should have read “4.1% absolute reduction.” In the Conflict of Interest Disclosures section, “Biosensors” was misspelled within Dr Roffi’s conflict of interest disclosure. The article has been corrected online.