

Letters

RESEARCH LETTER

Hospital Deaths in Patients With Sepsis From 2 Independent Cohorts

Sepsis, the inflammatory response to infection, affects millions of patients worldwide.¹ However, its effect on overall hospital mortality has not been measured. We quantified the contribution of sepsis to mortality in 2 complementary inpatient cohorts from Kaiser Permanente Northern California (KPNC) and the Healthcare Cost and Utilization Project Nationwide Inpatient Sample (NIS).

Methods | The KPNC cohort included 482 828 adults (aged ≥18 years) with overnight, nonobstetrical hospitalizations at 21 KPNC hospitals between 2010 and 2012.² Compared with all Northern California patients, KPNC patients have broadly similar health characteristics but higher income and educational attainment.³ The NIS, a nationally representative sample of 1051 hospitals, included 6.5 million unweighted adult hospitalizations in 2010.⁴

We used 2 approaches to identify patients with sepsis from *International Statistical Classification of Diseases, Ninth Revision, Clinical Modification* codes. The explicit approach identified those with codes 038 (septicemia), 995.91 (sepsis), 995.92 (severe sepsis), or 785.52 (septic shock). Because of the known underrecognition of sepsis, we also used an implicit approach adding patients with evidence of both infection and acute organ failure using the Angus implementation⁵ of sepsis consensus criteria. Within KPNC data, we delineated diagnoses when coded as present on admission, an important consideration for improving identification and treatment efforts. Furthermore, we linked 97.9% (n = 19 621) of all explicit sepsis cases present on admission in 2012 to KPNC quality improvement data, permitting stratification of patients by com-

mon sepsis severity criteria including early goal-directed therapy (EGDT) eligibility and serum lactate levels.⁶

In each cohort, we calculated the percentage of all inpatients admitted with sepsis, the sepsis hospital mortality rate, and the percentage and 95% confidence interval of hospital deaths occurring in patients with sepsis using Stata version 11.2 (StataCorp). The KPNC institutional review board approved the use of KPNC data with a waiver of informed consent and exempted NIS data from review.

Results | In the KPNC cohort, there were between 55 008 explicit (11.4% of total; 95% CI, 11.3%-11.5%) and 80 678 implicit (16.7%; 95% CI, 16.6%-16.8%) sepsis hospitalizations (Table 1); most occurrences of sepsis were present on admission. From the NIS cohort, 280 663 (4.3%; 95% CI, 4.3%-4.3%) hospitalizations met explicit sepsis criteria while 717 718 (10.9%; 95% CI, 10.9%-11.0%) met implicit criteria.

Of 14 206 KPNC inpatient deaths, 36.9% (95% CI, 36.1%-37.7%; explicit) to 55.9% (95% CI, 55.1%-56.7%; implicit) occurred among patients with sepsis, which was nearly all present on admission. Of 143 312 NIS deaths, 34.7% (95% CI, 34.4%-34.9%; explicit) to 52.0% (95% CI, 51.7%-52.2%; implicit) occurred among patients with sepsis. In the 2012 linked KPNC subset (Table 2), patients with sepsis meeting criteria for EGDT (n = 2536) comprised 32.6% (95% CI, 30.4%-34.7%) of sepsis deaths. In contrast, patients with sepsis, normal blood pressure, and measured lactate levels of less than 4 mmol/L (n = 15 095) comprised 55.9% (95% CI, 53.6%-58.1%) of sepsis deaths.

Discussion | In 2 complementary hospital cohorts, we found that sepsis contributed to 1 in every 2 to 3 deaths, and most of these patients had sepsis at admission. Given the prominent role it plays in hospital mortality, improved treatment of sepsis (po-

Table 1. Inpatients With Sepsis Diagnoses in the Kaiser Permanente Northern California Cohort and the Healthcare Cost and Utilization Project Nationwide Inpatient Sample^a

	Inpatients With Sepsis Diagnoses ^b				Nationwide Inpatient Sample (2010) (n = 1051 Hospitals) (143 312 Deaths/6 555 621 Admissions)	
	Kaiser Permanente Northern California (2010-2012) (n = 21 Hospitals) (14 206 Deaths/482 828 Admissions)				Explicit	Implicit
	Explicit	Explicit POA ^c	Implicit	Implicit POA ^c		
Hospitalizations	55 008 (11.4) [11.3-11.5]	50 520 (10.5) [10.4-10.5]	80 678 (16.7) [16.6-16.8]	73 933 (15.3) [15.2-15.4]	280 663 (4.3) [4.3-4.3]	717 718 (10.9) [10.9-11.0]
Hospital mortality	6272 (11.4) [11.1-11.7]	5238 (10.4) [10.1-10.6]	7941 (9.8) [9.6-10.0]	7391 (10.0) [9.8-10.2]	49 664 (17.7) [17.6-17.8]	74 451 (10.4) [10.3-10.4]
% (95% CI) of all hospital deaths among patients with sepsis	44.2 (43.3-45.0)	36.9 (36.1-37.7)	55.9 (55.1-56.7)	52.0 (51.2-52.8)	34.7 (34.4-34.9)	52.0 (51.7-52.2)

Abbreviation: POA, present on admission.

^a Case definitions are based on explicit sepsis diagnosis codes or the addition of cases identified using the Angus implementation⁵ of the International Consensus Conference Definition of Severe Sepsis (implicit).

^b Values expressed as No. (%) [95% CI] unless otherwise indicated.

^c Indicates diagnoses were POA.

Table 2. Hospital Mortality Among Patients With Sepsis Present on Admission^a

	Overall (n = 19 621)	Sepsis Severity Group ^b				
		Lactate ^c			Early Goal-Directed Therapy	
		Normal (n = 9067)	Intermediate (n = 6028)	None (n = 1990)	Yes (n = 1200)	No (n = 1336) ^d
Age, mean (SD), y	69 (17)	69 (17)	70 (16)	69 (17)	67 (16)	73 (16)
Laboratory and Acute Physiology Score, mean (SD) ^e	104 (40)	94 (33)	110 (39)	79 (43)	145 (39)	149 (39)
Direct admission to ICU, No. (%) [95% CI] ^f	3790 (19.3) [18.7-19.9]	879 (9.7) [9.1-10.3]	976 (16.2) [15.3-17.1]	267 (13.4) [11.9-14.9]	1087 (90.6) [88.9-92.2]	581 (43.5) [40.8-46.1]
Hospital deaths, No. (%) [95% CI]	1817 (9.3) [8.9-9.7]	477 (5.3) [4.8-5.7]	538 (8.9) [8.2-9.6]	211 (10.6) [9.2-12.0]	212 (17.7) [15.5-19.8]	379 (28.4) [25.9-30.8]
% (95% CI) of all sepsis deaths ^g		26.3 (24.2-28.3)	29.6 (27.5-31.7)	11.6 (10.1-13.1)	11.7 (10.2-13.1)	20.9 (19.0-22.7)

Abbreviation: ICU, intensive care unit.

^a Based on explicit diagnosis criteria and stratified by clinical characteristics using linked 2012 Kaiser Permanente Northern California quality improvement data. A total of 420 (2.1%) patients meeting explicit sepsis diagnosis criteria could not be matched with quality improvement data.

^b Patients were grouped sequentially into sepsis severity groups starting with those meeting standard criteria for early goal-directed therapy.

^c Remaining patients were stratified by lactate values: less than 2 mmol/L (normal), between 2 mmol/L or greater and less than 4 mmol/L

(intermediate), or not obtained (none).

^d These patients were eligible for early goal-directed therapy but they did not receive it.

^e Score range is from zero to a theoretical maximum of 414.²

^f Determined based on whether patients were transferred to a critical care hospital ward directly from the emergency department.

^g Row total may exceed 100% due to rounding.

tentially a final hospital pathway for multiple other underlying conditions) could offer meaningful improvements in population mortality.

Patients with initially less severe sepsis made up the majority of sepsis deaths. Performance improvement efforts in the treatment of sepsis have primarily focused on standardizing care for the most severely ill patients, whereas interventions for treating other patients with sepsis are less well defined. Given their prevalence, improving standardized care for patients with less severe sepsis could drive future reductions in hospital mortality.

Even though our findings were broadly consistent, the study's primary limitation results from potential inaccuracies and inconsistencies in case identification across cohorts. Prior strategies, based on administrative data, have demonstrated variability with respect to prevalence estimates and case accuracy, a factor that may have contributed to differences between cohorts in explicit sepsis mortality.⁵ Thus, we present granular data from the KPNC sepsis quality improvement program whose components include standardized case identification, manual chart validation, severity of illness risk adjustment, and treatment data; replication in other samples with similar granularity could be valuable.

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1. Levy MM, Fink MP, Marshall JC, et al; International Sepsis Definitions Conference. 2001 SCCM/ESICM/ACCP/ATS/SIS International Sepsis Definitions Conference. *Intensive Care Med*. 2003;29(4):530-538.

2. Escobar GJ, Gardner MN, Greene JD, Draper D, Kipnis P. Risk-adjusting hospital mortality using a comprehensive electronic record in an integrated health care delivery system. *Med Care*. 2013;51(5):446-453.

3. Gordon NP. Similarity of the adult Kaiser Permanente membership in Northern California to the insured and general population in Northern California: statistics from the 2009 California Health Interview Survey. http://www.dor.kaiser.org/external/chis_non_kp_2009. Accessed February 1, 2014.
4. Agency for Healthcare Research and Quality. Introduction to the HCUP Nationwide Inpatient Sample (NIS) 2010. http://www.hcup-us.ahrq.gov/db/nation/nis/NIS_Introduction_2010.jsp. Accessed February 1, 2014.
5. Iwashyna TJ, Odden A, Rohde J, et al. Identifying patients with severe sepsis using administrative claims: patient-level validation of the Angus implementation of the International Consensus Conference definition of severe sepsis [published online September 18, 2012]. *Med Care*. doi:10.1097/MLR.0b013e318268ac86.
6. Liu V, Morehouse JW, Soule J, Whippy A, Escobar GJ. Fluid volume, lactate values, and mortality in sepsis patients with intermediate lactate values. *Ann Am Thorac Soc*. 2013;10(5):466-473.

COMMENT & RESPONSE

Kidney Donation and Risk of ESRD

To the Editor Dr Muzaale and colleagues¹ matched all US kidney donors from 1994 to 2011 with a cohort of selected participants from the Third National Health and Nutrition Examination Survey (NHANES III) and found an increased cumulative incidence and lifetime risk of end-stage renal disease (ESRD) among the donors. The Editorial by Drs Gill and Tonelli² highlighted several methodological factors that may have biased the findings in the direction of finding increased risk to donors, and we identify an additional factor.

A review of eAppendix 1, which lists the exclusion criteria used for selecting healthy controls from the NHANES population, suggests that many individuals who might have been accepted as donors were excluded from the control group. These exclusions included any history of asthma, any cancer besides skin cancer, hypertension, and kidney stones. Selected patients with hypertension and kidney stones are not infrequently accepted as donors.^{3,4} We argue that a history of mild asthma or a remotely treated thyroid malignancy would not preclude donation. Similarly, functional limitations such as difficulty managing money or requiring special eating utensils were listed as exclusion criteria for NHANES controls, even though such limitations in isolation may not preclude live donation. These exclusions made the control group healthier and less likely to develop ESRD, thus biasing the findings in the direction of greater ESRD risk for donors.

We appreciate the difficulty in finding ideal controls in determining outcomes of live kidney donation, especially from retrospective data. Prior studies of kidney donors⁵ have been criticized as providing excessive reassurance about donor outcomes because the control populations did not exclude persons with diabetes and other individuals who would have been excluded from donation. The study by Muzaale et al¹ appears biased in the opposite direction, by excluding many individuals from the control group who might have been allowed to donate. As the authors pointed out, the magnitude of increased ESRD risk among donors was small—we suspect this risk is even smaller than reported and believe these findings should not discourage the many potential kidney donors for whom donation is safe.

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1. Muzaale AD, Massie AB, Wang M-C, et al. Risk of end-stage renal disease following live kidney donation. *JAMA*. 2014;311(6):579-586.
2. Gill JS, Tonelli M. Understanding rare adverse outcomes following living kidney donation. *JAMA*. 2014;311(6):577-578.
3. Mandelbrot DA, Pavlakis M, Danovitch GM, et al. The medical evaluation of living kidney donors: a survey of US transplant centers. *Am J Transplant*. 2007;7(10):2333-2343.
4. Davis CL, Cooper M. The state of US living kidney donors. *Clin J Am Soc Nephrol*. 2010;5(10):1873-1880.
5. Ibrahim HN, Foley R, Tan L, et al. Long-term consequences of kidney donation. *N Engl J Med*. 2009;360(5):459-469.

To the Editor The study by Dr Muzaale and colleagues,¹ analyzing United States data, reported a slightly increased rate of ESRD in kidney donors compared with selected healthy controls. This rate, however, was lower than that reported in the general population. A similar observation was recently reported from Norway.² Drs Gill and Tonelli³ have outlined some of the limitations of the analyses. We have additional concerns regarding the study design and conclusions.

First, compared with the healthy controls, the kidney donors (at time of donation) had a significantly higher systolic blood pressure and body mass index and were twice as likely to be smokers. Each of these factors is associated with increased risk of ESRD.

Second, the majority of cases of ESRD (83 of 99) occurred in related kidney donors, raising the possibility of a genetic contribution to these cases.⁴ Although the cause of ESRD was not provided in this study, in the Norwegian study, all cases of ESRD were in relatives and the majority were immunological, which would have affected both kidneys.²

Third, for the data analysis, 9364 controls were matched (using replacement matching) with 96 127 donors. How many times each control was used was not described. Given that this technique could magnify any differences, we wonder what the effect of using controls more than once was on the results. In terms of absolute numbers, 99 donors (0.1%) and 17 controls (0.2%) developed ESRD.

Last, in comparing lifetime risk of ESRD, the authors stated that the estimated lifetime risk for healthy controls younger than 30 years was 0 per 10 000. Others have reported much higher rates (2%-3% for whites; 7% for blacks).⁵ The comparison by Muzaale et al¹ may have lead to an overestimation of ESRD in donors.

Given these limitations, we are concerned that the analysis by Muzaale et al¹ was biased by confounding, and their assertion that “ESRD occurred in 23 white, 26 Hispanic, and 51 black individuals because they donated a kidney” is problematic. Additional long-term donor follow-up studies, addressing these limitations, should be done.

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