The Association Between Meningioma and Breast Cancer
A Systematic Review and Meta-analysis

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Abstract

IMPORTANCE A potential relationship between meningioma and breast cancer was suggested 70 years ago. However, to date, no conclusive evidence is available on this topic.

OBJECTIVE To provide a comprehensive review of the literature on the association of meningioma with breast cancer, supported by a meta-analysis.

DATA SOURCES A systematic PubMed search was performed up to April 2023 to identify articles on the association of meningioma with breast cancer. The following key words were used strategically: meningioma, breast cancer, breast carcinoma, association, relation.

STUDY SELECTION All studies reporting women diagnosed with meningioma and breast cancer were identified. The search strategy was not limited by study design or publication date but only included articles in English. Additional articles were identified via citation searching. Studies reporting a complete population of meningiomas or breast cancer patients throughout a specific study period and a proportion of patients with a second pathology could be used for the meta-analysis.

DATA EXTRACTION AND SYNTHESIS Data extraction was performed by 2 authors in accordance with the Preferred Reporting Items for Systematic Reviews (PRISMA) statement. Meta-analyses regarding both populations were performed using a random-effects model. Risk of bias was assessed.

MAIN OUTCOMES AND MEASURES The main measures were whether there was an increased prevalence of breast cancer in female patients with meningioma and whether there was an increased prevalence of meningioma in female patients with breast cancer.

RESULTS A total of 51 retrospective studies (case reports, case series, and cancer registry reports) describing 2238 patients with both diseases were identified; 18 studies qualified for prevalence analyses and meta-analysis. The random-effects meta-analysis (13 studies) revealed a significantly greater prevalence of breast cancer in female patients with meningioma than in the overall population (odds ratio [OR], 9.87; 95% CI, 7.31-13.32). Meningioma incidence in patients with breast cancer (11 studies) was greater than that in the baseline population; however, the difference according to the random-effects model was not statistically significant (OR, 1.41; 95% CI, 0.99-2.02).

CONCLUSIONS AND RELEVANCE This large systematic review and the meta-analysis on the association between meningioma and breast cancer found nearly 10-fold higher odds of breast cancer in female patients with meningioma compared with the general female population. These

(continued)
findings suggest that female patients with meningioma should be screened more intensively for breast cancer. Further research is required to identify the factors causing this association.

Introduction

The most common adult primary intracranial neoplasms are meningiomas arising from the meningothelial cells on the outer surface of the arachnoid villi.\(^1,2\) The age-adjusted incidence of meningiomas reported in the United States is 12.41 per 100,000 female individuals per year (2014-2018), and it continues to increase, according to the Central Brain Tumor Registry United States.\(^1\) The peak incidence of this typically benign tumor (ie, benign in 98% of patients; World Health Organization grade one, 80% of patients; grade two, 18.3% of patients; and grade three, 1.6% of patients) occurs from ages 50 to 70 years.\(^3,4\) In addition, this is a tumor with a clear female predisposition, occurring with a female to male ratio of 2.3 to 1.\(^1,4\)

Another typical female disease is breast cancer. With its continuously increasing incidence, this is the second most common type of cancer in female patients.\(^5\) Its worldwide incidence was 47.8 per 100,000 female individuals in 2020, according to the Global Cancer Observatory.\(^5\)

Seventy decades ago,\(^6\) a potential association between meningioma and breast cancer was first described. Since the initial case report by Lapresle et al\(^6\) in 1952 describing an individual with both illnesses, multiple case reports and a few larger patient series have suggested an association between meningioma and breast cancer. Schoenberg et al\(^7\) established this unique association in 1975, reporting 8 patients with breast cancer in a cancer registry of 402 patients with meningioma instead of the 3.4 expected patients.\(^7\)

More recently, studies have been conducted on a larger scale using cancer registries. However, these results demonstrate divergent conclusions. A 2020 study by Lopez-Rivera et al\(^8\) found an increased risk of meningioma after a breast cancer diagnosis.\(^8\) A 2016 study by Ji et al\(^9\) found a similar outcome in a portion of their investigated cohort. However, in 2013, Criscitiello et al\(^10\) reported no higher risk of subsequent meningioma in female patients with breast cancer. As a result, meningioma and breast cancer have not yet been conclusively linked.

Therefore, the aim of our study is to provide a comprehensive review of the published literature related to the association between meningioma and breast cancer, supported by prevalence analyses and a meta-analysis. We will present results considering both populations, specifically regarding the prevalence of meningioma among individuals with breast cancer, as well as the prevalence of breast cancer among individuals with meningioma.

Methods

Literature Search Strategy

This systematic review was performed up to April 2023 in accordance with the Preferred Reporting Items for Systematic Reviews and Meta-analyses (PRISMA) reporting guideline. All articles describing patients diagnosed with both intracranial or spinal meningioma and breast cancer or describing breast cancer and intracranial collision tumors were considered for review. Article selection was performed by a systematic search of PubMed to identify articles regarding an association between meningioma and breast cancer. The following key words were used: meningioma AND breast cancer OR breast carcinoma AND association OR relation. The search strategy was not limited by study design or publication date but only included key words in the English language; thus, at least the abstract had to be in English. Additional articles were identified via citation searching. The titles and abstracts of all records were screened by A.D. and V.D.M.
Regarding the first diagnosed tumor, histology, hormonal receptor expression, treatment received, age at diagnosis of the first tumor, and time between diagnoses of the 2 diseases were collected from full texts when available. Data extraction was performed by A.D. and V.D.M.

**Selection Criteria**

Subsequently, a selection was made of articles for which odds ratio (OR) calculations and prevalence analysis could be carried out on the basis of the following inclusion and exclusion criteria: studies were only included in the analysis if a complete population of patients with meningioma or breast cancer was presented throughout a specific study period, together with a proportion of patients with a second condition, whether breast cancer or meningioma. Consequently, case reports and larger case series lacking a total population were excluded. Studies with overlapping data from the same cancer registry were excluded.

**Statistical Analysis**

**Prevalence Analysis**

After selecting articles suitable for the analysis, articles were classified into 2 groups. The first group consisted of studies starting from a population of female patients with meningioma in which the number of patients who had breast cancer also was examined. Breast cancer may have been diagnosed before or after meningioma. The second group consisted of studies starting with female patients with breast cancer and involved examining the number of female patients with meningiomas regardless of whether the diagnosis of meningioma was made before or after the diagnosis of breast cancer during the study period. Studies performing bidirectional analyses were classified into each of these 2 groups. Few studies indicated which diagnosis was first. Due to a lack of data regarding the first diagnosed tumor, histology, hormonal receptor expression, treatment received, age at diagnosis of the first tumor, and time between diagnoses of the 2 diseases in most of the literature reviewed, no subanalyses could be conducted.

The proportions of patients with the second condition were calculated for each study. For group 1, the proportion of patients with breast cancer per study (prevalence per study) in the meningioma group was determined. This was then compared with the 2020 breast cancer prevalence, adjusted for the country where the study was conducted. Using MedCalc version 20.111 (MedCalc Software), a 1-proportion test was run to determine whether the population under study deviated significantly from the general prevalence. The proportion of patients with meningioma per study (prevalence per study) in the breast cancer group was determined. This was compared with the 2020 meningioma prevalence in Belgium. Identical procedures were followed for group 2.

**Meta-analysis**

Based on these results, a meta-analysis was carried out, gathering results using MedCalc, SPSS Statistics version 28 (IBM), and NCSS version 10 (NCSS Statistical Software). Heterogeneity between studies was evaluated using Cochran Q statistics and the I² statistic. A significant Q statistic value suggests that the homogeneity hypothesis of the association set can be rejected, while an I² value greater than 50% indicates strong heterogeneity. A restricted maximum likelihood (REML) random-effects model was used. ORs were considered. Egger and Peter tests were used to detect possible publication bias. Meta-analyses were conducted for each of the 2 groups. The Meta-analyses of Observational Studies in Epidemiology (MOOSE) checklist was used as a guide.

**Results**

**Systematic Review**

**Figure 1** shows the flowchart for the study selection process at different stages. A total of 306 articles were identified: 261 records were found by searching the PubMed database, and 45 studies were found manually and cross-referenced with identified articles’ corresponding reference lists.
English-language full text was not available for 26 records; hence, they were excluded. Finally, 53 studies\textsuperscript{6-57} were included for this systematic review: 27 case reports,\textsuperscript{6,13-36} 14 case series,\textsuperscript{10,37-48} and 12 cancer registry studies.\textsuperscript{7,9,49-57} The case reports, case series, and cancer registry studies described a total of 2238 individual patients diagnosed with both pathologies, including patients with intracranial collision tumors.

### Study Selection

Twenty-three articles were selected as potentially suitable for the analysis of odds, ORs, and prevalence. After the exclusion of 5 studies\textsuperscript{43,49,53,54,56} for the reasons described in the Table, 18 studies\textsuperscript{6-10,37,39,41,42,44,45,47,48,50-52,55,57} were ultimately included for these analyses. The study by Maiuri et al\textsuperscript{43} found a particularly low value for the breast cancer OR.\textsuperscript{43} We observed an underreporting bias in this instance. Of 660 patients with meningioma who underwent surgery, 3 had breast cancer and brain metastases. Maiuri et al\textsuperscript{43} focused solely on the number of patients with breast cancer who had brain metastases. Patients with nonmetastatic breast cancer were not discussed in the study by Maiuri et al\textsuperscript{43}; therefore, this study was excluded from the subsequent analyses.

Eight studies\textsuperscript{6,7,41,42,45,48,52,57} analyzed a population of patients with meningioma, 5 studies\textsuperscript{6,10,37,39,44} analyzed a population of patients with breast cancer, and the other 5 studies\textsuperscript{39,47,50,51,55} conducted a bidirectional analysis. The study by Ji et al\textsuperscript{9} separated analyses into 2 sections, distinguishing between a cohort of patients diagnosed with breast cancer prior to 1987 and a cohort diagnosed after 1987. The latter group was characterized as having been exposed to tamoxifen. Consequently, these 2 components of the study were regarded as independent units in the analysis.

This yielded a total of 13 analyses based on the meningioma population and 11 analyses based on the breast cancer population. As only retrospective studies and case series were available, the level of evidence of the included studies was 3 or 4, according to the Oxford Centre for Evidence-based Medicine for ratings of individual studies.

Using the incidence ratios per year, the number of observed cases can be compared with the number of expected cases, and thus, the standardized incidence ratio (SIR) can be calculated. Of 13 studies\textsuperscript{6,7,39,41,42,45,47,48,50-52,55,57} assessed for the meningioma population, 6 expressed the risk of developing breast cancer as SIR or similar ratios. The most recent study, conducted by Ben Lassan et al,\textsuperscript{57} reported an SIR of 1.6 (95% CI, 1.4-1.9) for invasive breast cancer and 1.7 (95% CI, 1.2-2.6) for in situ breast cancer in their cohort of female individuals diagnosed with meningioma. A study by Rao...
et al\textsuperscript{55} reported that the cumulative observed rate of breast cancer in female individuals previously diagnosed with meningioma was 58 times the expected rate.\textsuperscript{55} A study by Custer et al\textsuperscript{50} found that the risk of breast cancer was not significantly increased after a diagnosis of meningioma in their cohort (SIR, 1.54; 95% CI, 0.77-2.75).\textsuperscript{50} A study by Helseth et al\textsuperscript{51} identified a ratio of observed vs expected breast cancer of 21 to 11.9 (SIR, 1.75; 95% CI, 1.08-2.68) with regard to developing a subsequent breast cancer.\textsuperscript{51} Two studies by Jacobs et al\textsuperscript{52,53} reported an SIR of 1.43 (95% CI, 0.74-2.50), regardless of whether the breast cancer diagnosis manifested before or after the meningioma diagnosis. Last, the 1975 study by Schoenberg et al\textsuperscript{7} was the first to identify a statistically significantly increased risk of breast cancer (SIR, 2.37) in the presence of meningiomas in female patients.\textsuperscript{7}

Of 11 studies\textsuperscript{8,10,37,39,44,47,50,51,55} assessed for the breast cancer population, 6 expressed the risk of developing meningioma as SIR or similar. The study by Helseth et al\textsuperscript{51} reported an SIR of 1.54 (95% CI, 0.97-2.34), Custer et al\textsuperscript{50} reported an SIR of 1.40 (95% CI, 0.67-2.58), Rao et al\textsuperscript{55} found 80 (95% CI, 72-89), Ji et al\textsuperscript{5} reported an SIR of 1.53 (95% CI, 1.30-1.81) without tamoxifen exposure and 1.06 (95% CI, 0.84-1.32) with tamoxifen exposure, and Lopez-Rivera et al\textsuperscript{8} discovered an SIR of 1.26 (95% CI, 1.19-1.33) within the first 5 years after the diagnosis of breast cancer. According to Lopez-Rivera et al,\textsuperscript{8} patients with breast cancer aged 18 to 49 years (SIR, 2.16; 95% CI, 1.78-2.61) and those with a more advanced tumor stage (stage IV: SIR, 2.39; 95% CI, 1.71-3.25; \( P = .05 \)) were at a higher risk of developing meningioma. Patients in subgroups according to hormone receptor expression and treatment modality were at comparable risk vs the overall population.\textsuperscript{8} A study by Criscitiello et al\textsuperscript{10} found no statistically significant increased risk of developing meningioma.

**Meta-analysis**

**Meningioma Population**

When comparing the studies,\textsuperscript{6,7,39,41,42,45,47,48,50-52,55,57} the heterogeneity test results were highly statistically significant (Cochran \( Q = 92.9; \) \( df = 12; \) \( P < .001 \)) with an \( I^2 \) of 93%, indicating that the between-study heterogeneity was high. Therefore, a random-effects model was chosen for the meta-analysis. Possible publication bias was detected (Egger linear regression intercept [SE], 2.57 [0.36]; \( P < .001 \); Peter linear regression intercept [SE], 2.12 [0.24]; \( P < .001 \)). The meta-analysis, based on the random-effects model, revealed a considerably greater prevalence of breast cancer in female patients with meningioma compared with the baseline prevalence of breast cancer in the general population (OR, 9.87; 95% CI, 7.31-13.32). **Figure 2** illustrates the results of the meta-analysis.

**Breast Cancer Population**

When comparing the breast cancer studies,\textsuperscript{8,10,37,39,44,47,50,51,55} the heterogeneity test results were highly statistically significant (Cochran \( Q = 345.5; \) \( df = 10; \) \( P < .001 \)), with an \( I^2 \) of 98%, indicating that the between-study heterogeneity was high. Therefore, a random-effects model was chosen for the meta-analysis. No publication bias was observed using Egger test (linear regression intercept [SE], 0.25 [0.40]; \( P = .56 \)) nor Peter test (linear regression intercept [SE], 0.38 [0.28]; \( P = .22 \)). The random-effects model did not reach statistical significance (OR, 1.41; 95% CI, 0.99-2.02). **Figure 3** illustrates the results of the meta-analysis.

**Table. List of Excluded Articles**

<table>
<thead>
<tr>
<th>Publication</th>
<th>Reason for exclusion</th>
<th>Year</th>
</tr>
</thead>
<tbody>
<tr>
<td>Adami et al,\textsuperscript{49} 1984</td>
<td>Overlap of data with Ji et al,\textsuperscript{9} 2016</td>
<td></td>
</tr>
<tr>
<td>Jacobs et al,\textsuperscript{13} 1992</td>
<td>Overlap of data with Jacobs et al,\textsuperscript{12} 1987</td>
<td></td>
</tr>
<tr>
<td>Malmer et al,\textsuperscript{56} 2000</td>
<td>Overlap of data with Ji et al,\textsuperscript{9} 2016</td>
<td></td>
</tr>
<tr>
<td>Maiuri et al,\textsuperscript{43} 2002</td>
<td>Reporting bias</td>
<td></td>
</tr>
<tr>
<td>Milano et al,\textsuperscript{34} 2014</td>
<td>Overlap of data with Lopez-Rivera et al,\textsuperscript{8} 2020</td>
<td></td>
</tr>
</tbody>
</table>
Discussion

To our knowledge, this systematic review and meta-analysis is both the largest systematic review and the first meta-analysis that has been conducted on the association between meningioma and breast cancer. The main objective of the study was to clarify whether there was an increased prevalence of breast cancer among female patients with meningioma, and vice versa. Based on the meta-analysis of 18 retrospective studies, 6-10, 37, 39, 41, 42, 44, 45, 47, 48, 50-52, 55, 57 there was an association between these diseases.

In female patients with meningioma, we found approximately 10-fold higher odds of developing breast cancer during their lifetime based on 13 cohorts. 6,7,39,41,42,45,47,48,50-52,55,57 In previous literature, the risk of having breast cancer while diagnosed with a meningioma was explored using state and national cancer registries. To our knowledge, only 3 cohorts have previously reported a statistically significant increase in the risk of developing breast cancer in female patients with meningiomas. 7,53,55

In female patients with breast cancer, we found that odds of developing meningioma during the lifetime were increased by 1.42-fold, based on 11 cohorts. There was no statistically significantly increased prevalence of meningioma in the population of female patients with breast cancer compared with the control population. In 5 studies, 8,9,50,51,55 an SIR for developing a subsequent meningioma after breast cancer diagnosis in females was reported. Only 3 previous cohorts 8,9,55

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### Figure 2. Estimated Odds Ratios (ORs) for Breast Cancer in Each Meningioma Study and Overall

<table>
<thead>
<tr>
<th>Study</th>
<th>OR (95% CI)</th>
<th>Weight, %</th>
</tr>
</thead>
<tbody>
<tr>
<td>Lapresle et al, 6 1952</td>
<td>1.80 (0.25-12.96)</td>
<td>1.95</td>
</tr>
<tr>
<td>Schoenberg et al, 7 1975</td>
<td>3.15 (1.56-6.35)</td>
<td>7.40</td>
</tr>
<tr>
<td>Smith-Behn et al, 42 1992</td>
<td>3.79 (0.52-27.52)</td>
<td>1.93</td>
</tr>
<tr>
<td>Salvati et al, 45 1996</td>
<td>4.89 (2.52-9.51)</td>
<td>7.70</td>
</tr>
<tr>
<td>Custer et al, 50 2002</td>
<td>6.49 (4.31-9.77)</td>
<td>10.06</td>
</tr>
<tr>
<td>Ben Lassan et al, 27 2022</td>
<td>8.04 (6.92-9.35)</td>
<td>12.00</td>
</tr>
<tr>
<td>Helseth et al, 21 1989</td>
<td>8.30 (6.10-11.28)</td>
<td>10.95</td>
</tr>
<tr>
<td>Jacob et al, 22 1987</td>
<td>11.29 (6.28-20.29)</td>
<td>8.41</td>
</tr>
<tr>
<td>Lieu et al, 55 2003</td>
<td>11.62 (4.32-31.26)</td>
<td>5.28</td>
</tr>
<tr>
<td>Kuro et al, 51 2001</td>
<td>17.81 (8.21-38.62)</td>
<td>6.79</td>
</tr>
<tr>
<td>Wang et al, 48 2022</td>
<td>29.30 (14.34-59.88)</td>
<td>7.27</td>
</tr>
<tr>
<td>Bonito et al, 39 1993</td>
<td>31.86 (17.00-59.71)</td>
<td>8.03</td>
</tr>
<tr>
<td>Mean</td>
<td>9.87 (7.31-13.32)</td>
<td>NA</td>
</tr>
</tbody>
</table>

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### Figure 3. Estimated Odds Ratios (ORs) for Meningioma in Each Breast Cancer Study and Overall

<table>
<thead>
<tr>
<th>Study</th>
<th>OR (95% CI)</th>
<th>Weight, %</th>
</tr>
</thead>
<tbody>
<tr>
<td>Ji et al, 9 2016 (cohort 2)</td>
<td>0.43 (0.35-0.54)</td>
<td>10.96</td>
</tr>
<tr>
<td>Smith-Behn et al, 47 1992</td>
<td>0.70 (0.40-4.39)</td>
<td>2.51</td>
</tr>
<tr>
<td>Helseth et al, 21 1989</td>
<td>0.72 (0.54-0.98)</td>
<td>10.59</td>
</tr>
<tr>
<td>Custer et al, 56 2002</td>
<td>0.80 (0.53-1.19)</td>
<td>9.99</td>
</tr>
<tr>
<td>Burns et al, 31 1986</td>
<td>0.89 (0.37-2.15)</td>
<td>6.72</td>
</tr>
<tr>
<td>Ji et al, 2 2016 (cohort 1)</td>
<td>1.05 (0.89-1.24)</td>
<td>11.18</td>
</tr>
<tr>
<td>Lopez-Rivera et al, 3 2020</td>
<td>1.47 (1.38-1.56)</td>
<td>11.42</td>
</tr>
<tr>
<td>Criscitiello et al, 38 2013</td>
<td>1.91 (1.36-2.77)</td>
<td>10.35</td>
</tr>
<tr>
<td>Rao et al, 55 2009</td>
<td>2.51 (2.31-2.74)</td>
<td>11.38</td>
</tr>
<tr>
<td>Bonito et al, 39 1993</td>
<td>6.11 (4.36-10.79)</td>
<td>8.84</td>
</tr>
<tr>
<td>Mehta et al, 44 1983</td>
<td>8.92 (3.33-23.91)</td>
<td>6.06</td>
</tr>
<tr>
<td>Mean</td>
<td>1.42 (1.00-2.02)</td>
<td>NA</td>
</tr>
</tbody>
</table>

Dots indicate estimates; size of dots, weight; whiskers, 95% CIs; diamond, overall association estimated using a random-effects model; NA, not applicable.
reported a statistically significant increase in the likelihood of developing meningioma in female patients with breast cancer.

Future Perspective

Previous study groups have proposed possible explanations for the association between meningioma and breast cancer, such as hormone receptor expression in both diseases or adverse effects of radiation therapy and hormonal therapy. Moreover, the role of overactivation of the MYC oncogene has been hypothesized.18,38

Both types of tumors have been linked to hormone receptors in their genesis and progression. Approximately 88% of meningiomas express progesterone receptors (PR), but only 30% express estrogen receptors (ER).58,59 The ER is expressed in 70% of breast cancers; the PR is expressed in 50% of ER-positive breast cancers but is rarely seen in ER-negative breast cancers.60-63 However, the studies by Criscitiello et al10 and Lopez-Rivera et al8 did not find any significant association between the hormonal receptor status of breast cancer and the risk of meningioma.

Ionizing radiation is a well-known risk factor for meningioma and is frequently used as a treatment for breast cancer.64-66 This might be a contributing factor to the cooccurrence of both diseases. However, the study by Lopez-Rivera et al8 failed to show a difference in the observed risk of meningioma among patients treated with vs without radiotherapy for breast cancer. Hormone therapy, particularly synthetic progestagens, has been associated with an increased risk of both meningioma and breast cancer.58,67,68

Undoubtedly, further research is required to unravel the potential risk factors related to the cooccurrence of benign meningioma and malignant breast cancer, and the findings from this meta-analysis should be considered for implementation in guidelines regarding screening for these 2 diseases in female patients, specifically screening for breast cancer in female patients with meningioma.

Limitations

This study has several limitations. All previous studies have expressed the occurrence of both diseases in terms of incidence per year. Incidence rates are more widely used than rates of prevalence to express the prevalence of a disease. Incidence baseline values in the population per year and country are widely available. However, performing a meta-analysis on the basis of incidence rates was not possible, as we did not have information in all studies about which diagnosis was made first and at what point in the study the diagnoses were established; therefore, we could not correct for the number of study years. For these reasons, we opted to use prevalence. At the end of the study, the total number of individuals with both meningioma and breast cancer in the total population was assessed. These prevalences were compared with baseline prevalences. However, data on prevalence were scarce. We used a baseline prevalence measured in Belgium in 2020. For breast cancer, we also used prevalence rates from 2020, but we had sufficient breast cancer data to adjust the values for each country where the study was performed. Older, reliable prevalence rates were not available in the literature. As the rates of incidence of both meningiomas and breast cancer are increasing with time, we most likely overestimated them in our meta-analyses; this fact further strengthens our conclusions. Among studies using baseline prevalence of meningiomas, 45% date from before 2000. Baseline prevalence was based on histologically confirmed meningiomas, and in some of the studies, meningioma was diagnosed based on radiological findings without histologic confirmation.

On the other hand, the prevalence of breast cancer in the global population used for our study was overestimated, since 2020 numbers were used and adjusted per country. 54% date from before 2000. Since the incidence of breast cancer is increasing each year, the study was carried out under the least favorable conditions, and it must be concluded that the OR and relative risk are potentially greater than those estimated in our work. On the other hand, we were unable to adjust for age, race, and diagnosis year when calculating the OR and relative risk due to a lack of data.
A limitation of our study could be linked to a possible publication bias in the meta-analysis of group 1. Another limitation of the study is that we were unable to extract sufficient data from the investigated studies regarding the first diagnosed tumor, histology, hormonal receptor expression, treatment received, age at diagnosis of the first tumor, and time between diagnoses of the 2 diseases to conduct subanalyses.

In addition, we would also like to report that 1 ad hoc exclusion criteria was used due to a reporting bias in 1 study (Maiuri et al²⁴), discovered later in the analysis process. This also constitutes a limitation of the study.

Conclusions

To our knowledge, this study is the largest systematic review and the first meta-analysis on the association between meningioma and breast cancer, and we found that female patients diagnosed with a meningioma had nearly 10-fold higher odds of breast cancer compared with the general female population. Based on this finding, female patients with meningioma should be screened more intensively for breast cancer.


SUPPLEMENT.

Data Sharing Statement